Estimation subject to reporting delays and incomplete event adjudication with applications to disability insurance

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Overview

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Problem outline

- **Goal**: Estimate individual pricing and reserving models for disability insurance products.
 - Lack of steady-state makes aggregate reserves problematic.
- Represent disability insurance schemes using multistate model.
 - Capture a priori known structure of payments and model intertemporal dependencies.
- Hazard rates characterize conditional distribution (needed for reserving) and can be estimated with censored data.
- **Problem**: Biased sampling due to reporting delays (IBNR) and incomplete event adjudication (RBNS).

 \Rightarrow Fitting model directly to observed data leads to severe bias!

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Where we are going...



Figure 1: Fitted rates (lines) and occurrence-exposure rates (points) for the proposed method (black) and the naive method (gray). Disability rates are shown on the left and reactivation rates on the right.

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Classic multi-state model

- State process $\{Y(s)\}_{s\geq 0}$: When insurance events occur.
- State space for Y:



• Counting process representation $N_{jk}(t) = \#\{s \le t : Y(s-) = j, Y(s) = k\}.$

• Marked point process representation $(T_m, Y_m)_{m \ge 1}$.

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Classic multi-state estimation

• Statistical model: Parametric intensity of $N_{jk}(s)$:

$$s \mapsto \mu_{jk}(s, \theta), \quad \theta \in \mathbb{R}^d.$$

- Maximum likelihood estimator (MLE) with discretization $0 = t_0 < t_1 < \cdots < t_l = t$:
 - Occurrences and exposures:

$$egin{aligned} & O_{jk}(t_i) = N_{jk}(t_{i+1}) - N_{jk}(t_i), \ & E_j(t_i) = \int_{t_i}^{t_{i+1}} \mathbbm{1}\{Y(s) = j\} \, \mathrm{d}s. \end{aligned}$$

 Input (O_{jk}(t_i))_{j,k,i} as independent Poisson observations with mean (μ_{jk}(t_i, θ)E_j(t_i))_{j,k,i} (see Lindsey (1995)).

Research problem

Problem: $\{Y(s)\}_{s \le t}$ not available at time t due to reporting and processing delays \Rightarrow MLE cannot be used directly.

Focus of presentation:

- Illustration of the problem.
- Outline of mathematical approach & literature.
- Overview of results.
- Data applications.

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Classic multi-state model

Main example: Disability insurance with reactivation.



Figure 2: State space for state process Y.

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How to estimate?

- **Question**: How to estimate θ for $\mu_{jk}(s, \theta)$ based on observed information?
- Naive approach: Use old data (backcensoring).
- **Our approach**: Derive estimators $\hat{\theta}$ under IBNR and RBNS contaminated data for multi-state models.

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Pros and cons over naive approach

Advantages

- More efficient use of data \Rightarrow less estimation risk.
- Use new data faster \Rightarrow capture new trends.
- Estimates of IBNR and RBNS applies to reserving.

Disadvantages

- Additional model elements \Rightarrow added estimation and model risk.
- Requires detailed data.
- Slightly more complicated to implement.

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Contributions

Contributions:

- Handle reporting delays for general multi-state model.
 - Insurance literature: Antonio & Plat (2014) and Bücher & Rosenstock (2024) for Marked Poisson process; Badescu et al. (2016,2019) for Marked Cox process.
- Handle incomplete event adjudication for hazard estimation, general multi-state model, and dynamical conditioning in adjudication probabilities.
 - Literature: Cook & Kosorok (2004), Bladt & Furrer (2023).
- Simultaneous treatment of reporting delays and incomplete event adjudication.
- Large-sample properties of the estimators.



Estimator construction (simplified)

- First estimate conditional reporting delay and adjudication probabilities.
- Estimator $\hat{\theta}$ can then (approximately) be based on observed occurrences $O_{jk}^{obs}(t_i)$ and exposures $E_j^{obs}(t_i)$ after modifying as follows:

$$egin{aligned} & O_{jk}(t_i) \leftarrow O_{jk}^{ ext{obs}}(t_i) imes \hat{P}(ext{Confirm } O_{jk}^{ ext{obs}}(t_i) \mid \mathcal{F}_t^{ ext{obs}}), \ & E_{jk}(t_i) \leftarrow E_j^{ ext{obs}}(t_i) imes \hat{P}(ext{Reporting delay} < t - t_i \mid \{Y(s)\}_{s < t_i} \ &, Y(t_i) = k). \end{aligned}$$

• Large-sample properties: Consistency, asymptotic normality, and Efron's simple nonparametric bootstrap is valid.

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Simulation study

400 samples of size n = 1500 with t = 5.



Figure 3: Event history model (left) and adjudication model (right). Symbols U and ξ indicate the presence of reporting delays and adjudication processes, respectively.

Setup:

- Moderately large transition rates.
- Reporting delays with mean 1.
- Confirm 40% of jumps.

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Simulation results

	Propos	ed method	Ora	cle	Approx	kimation	Na	ive
Parameter	Bias	SD	Bias	SD	Bias	SD	Bias	SD
$\theta_1 = \log 0.15$	004	.067	008	.031	010	.067	010	.066
$\theta_2 = 0.1$	000	.020	001	.020	006	.020	006	.021
$\theta_3 = 0.4$.003	.078	.003	.078	002	.078	000	.079
$\theta_4 = \log 0.1$.003	.084	.001	.083	.012	.091	051	.082
$\theta_5 = 0.03$.000	.012	000	.013	006	.016	015	.014
$\theta_6 = -0.3$	000	.094	001	.088	.007	.094	007	.090
$\theta_7 = -0.3$	011	.066	011	.054	012	.066	.148	.069

Table 1: Bias and empirical standard deviation (SD) of the estimator $\hat{\theta}_n$ based on 400 simulations of size n = 1500.

Overall:

- Bias: Oracle=Proposed method<Approximation≪Naive.
- SD: Oracle<Proposed method=Approximation=Naive.

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Application to real data: Model

Disability insurance data.

- Disability exposure and occurrences.
- Reactivation exposure and occurrences.
- Disability reporting delays.
- Adjudications.

Time window [0, *t*] is [31/01/2015, 01/09/2019].



Figure 4: Event history model (left) and adjudication model (right). For events, active is *a*, disabled is *i*, reactivated is *r*, and dead is *d*. For adjudications, active report is 1, inactive report is 2, adjudicated is 3, and dead is 4.



Application to real data: Results



Figure 5: Fitted rates (lines) and occurrence-exposure rates (points) for the proposed method (black) and the naive method (gray). Disability rates are shown on the left and reactivation rates on the right.

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Related paper: Reserving



Figure 6: Portfolio level reserve decomposed by category.

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Thank you for your attention!

